u^{b}

b UNIVERSITÄT BERN

Faculty of Business, Economics and Social Sciences

Department of Economics

Life-Cycle Inequality: the Black and White Differential

Giacomo De Giorgi, Luca Gambetti, Costanza Naguib

23-01

January, 2023

DISCUSSION PAPERS

Schanzeneckstrasse 1 CH-3012 Bern, Switzerland http://www.vwi.unibe.ch

Life-Cycle Inequality: the Black and White Differential*

Giacomo De Giorgi[†] GSEM-University of Geneva BREAD, CEPR, IPA Luca Gambetti[‡] Universitat Autònoma de Barcelona, BSE Università di Torino, and Collegio Carlo Alberto

Costanza Naguib[§] University of Bern and GSEM

Abstract

With 20 years of PSID data, we document persistent racial differentials in life-cycle consumption dynamics. Starting from similar positions in the consumption distribution Blacks end up in lower percentiles than Whites. Education, income, and wealth are three key drivers of these different dynamics. Blacks tend to save less, and hence have less buffer than the Whites to prevent them from falling in the lower part of the consumption distribution.

JEL classification: E21, D12, C3

Keywords: Consumption, Income, Savings, Inequality, Persistence, Life-Cycle.

^{*}We would like to thank Costas Meghir, Robert Moffitt, and Luigi Pistaferri for the insightful discussions. Thanks to workshop participants at the First Dondena Workshop, and the BGSE Summer Forum on Inequality, and seminar participants at the University of Warwick, and the University of Zurich.

[†]Contact: Giacomo De Giorgi. E-mail: giacomo.degiorgi@gmail.com. De Giorgi gratefully acknowledges the financial support of the EU, through the Marie Curie CIG grant FP7-631510, and the SNF no. 100018-182243.

[‡]Luca Gambetti acknowledges the financial support from the Spanish Ministry of Science and Innovation, through the Severo Ochoa Programme for Centres of Excellence in R&D (CEX2019-000915-S), the financial support of the Spanish Ministry of Science, Innovation and Universities through grant PGC2018-094364-B-I00, and the Barcelona Graduate School Research Network.

[§]Contact: Costanza Naguib. E-mail: costanza.naguib@vwi.unibe.ch

1 Introduction

Economic inequality remains one of the major challenges for policy-makers and economists. In recent years there has been a burgeoning literature in this area. Starting from Piketty and Saez [2003], several authors have documented and investigated the rising income, and wage inequality, in the US and other countries (see Autor et al. [2008], Bonhomme and Robin [2009a], Primiceri and Van Rens [2009], Heathcote et al. [2010], Atkinson et al. [2011], Auten et al. [2013], Attanasio and Pistaferri [2014, 2016], Blundell [2014], Chetty et al. [2014a]). Given that consumption is tightly related to permanent income and it is a driver of individual utility more than income itself, a related literature has also stressed the importance of focusing on consumption inequality in order to draw conclusions about households' well-being (see for example Blundell and Preston [1998], Meyer and Sullivan [2003], Krueger and Perri [2006], Blundell et al. [2008], Attanasio et al. [2014], Aguiar and Bils [2015], Attanasio and Pistaferri [2016], Blundell et al. [2016]).

A stream of the literature, close in spirit to the current paper, has investigated the existing differences in earnings levels between Black and White individuals in the US. Chetty et al. [2014b] study the influence of race on intergenerational income mobility via the channel of geographical segregation. Blau and Beller [1992] put under scrutiny gender differences in the Black-White earnings gap, whereas the quantity and quality of schooling has also been found to be a relevant driver of the black/white earnings gap (see e.g. Card and Krueger [1992], Heckman et al. [2000] and Bayer and Charles [2018]) Further, the college wage premium appears different between black and white workers (Chay and Lee [2000]). Black workers are less often in performance-pay jobs, and this exacerbates the racial earnings gap (Heywood and Parent [2012]). Oaxaca and Ransom [1994] find evidence of discrimination in the labor market against Black workers, and Peoples and Talley [2001] show that privatization, e.g. in the public transport sector, is associated with a decline in the racial earnings differential. Further, civil right legislation helped reducing the Black/White earnings gap in the 1960's (Card and Krueger [1993]).

The current paper focuses on a relatively unexplored area: the differential in consumption life-cycle dynamics between Blacks and Whites in the US over the past two decades. To the best of our knowledge, apart from the work by Charles et al. [2009], who show that Blacks devote a larger share of their expenditure bundles to *conspicuous* goods than do comparable Whites, racial differences in consumption, and in particular in consumption persistence over the life cycle, have not been investigated in the literature.

We use PSID data from 1999 to 2017 to document persistent racial differentials in life-cycle consumption dynamics across the distribution. More specifically, we document large differences between Blacks and Whites in terms of mobility along the consumption distribution. Blacks, independently of the initial percentiles, tend, over the life-cycle, to end up in lower percentiles than white individuals. Individual characteristics like age and gender do not seem to account for these differences, while education, income and wealth, when considered together, essentially close the gap at the top and substantially reduce it at the bottom of the distribution.

Our paper is close to the work of Chetty et al. [2020] on the inter-generational persistence of racial differences in income. In particular, the authors focus on the "intergenerational gaps", i.e. differences by race in children's incomes conditional on parental income. They find that such gaps are persistent over time and, conditional on parent income, they are mainly driven by differences in wages and employment rates. Further, Ganong et al. [2020] use bank data matched with voter registry and firm-wide wage changes data in order to estimate the transmission of unexpected income shocks into consumption by race. They, too, link race differentials in the degree of insurance against shocks to the different amount of wealth held by Blacks and Whites. However, differently from our work, they exclusively focus on short-term consumption changes in reaction to income shocks. Our paper focuses instead on life-cycle racial differences in consumption persistence.

The remainder of the paper is organized as follows: Section 2 presents the data. Section 3 provides evidence on the (unconditional) consumption persistence for Blacks and Whites. Section 4 discuss the main drivers of Black/White differentials in consumption. Section 5 explores conditional consumption persistence of Blacks and Whites. Section 6 concludes. In Appendix A we cover some additional analyses on income and consumption life-cycle profiles, while in Appendix B we present a series of robustness checks.

2 Data

We use data from the Panel Study of Income Dynamics (PSID), a longitudinal survey conducted by the University of Michigan. The PSID is a nationally representative sample of households. The PSID collects a wide range of variables, including information on demographics, income, and consumption. Most data is collected at the household level, though information for PSID-defined household "heads" and "wives" is also gathered. A limited selection of questions are asked about other family members. Typically, a family head is the male in a married pair with primary financial responsibility for the family. A wife is the female counterpart of the married couple. Females only qualify as heads in single adult households (single males can also be heads, of course). If a female head marries a man, he becomes the new head and the woman's classification changes to 'wife.'

2.1 Building the Dataset

To create our dataset, we append together all waves from 1999-2017. This choice is dictated by the fact that information on actual consumption is only collected starting in 1999.¹ We include only current heads, since they are the individuals with the richest and most consistent set of observables over time. As there is one head per household, our analysis is therefore effectively at the household level.

We then create a consistent race indicator for all individuals. The PSID asked heads to identify their race in every wave. For all heads, we assign race as the mode value of race from all reported years. Due to the limited sample size of some reported races, we only keep individuals identifying themselves as Black or White. Our full sample, for the years under scrutiny, includes 153,592 individual-year observations.

2.2 Household Income

The PSID consistently asks respondents to report their household's total monetary income, defined as the sum of the taxable income of the head and wife, the total transfers of the head and the wife, the taxable income of other family unit members, and the transfer income of other family unit members. Beginning with the 1994 wave, the measure also includes total family Social Security income.

Any negative or zero values are recorded to \$1 in the PSID, and because this practice occurs for many years, we apply the same rule to the remaining years of data. To convert nominal incomes to real terms, we divide the nominal measure by the Consumer Price Index (CPI). In order to create a per capita measure, we then divide total family income by an Adult Equivalent scale, given by:

$$AE = 1 + 0.7(A - 1) + 0.5K \tag{1}$$

where A is the number of adults in the household and K is the number of children in the household. This scale assigns a value of 1 to the first household member, of 0.7 to each other adult in the household and 0.5 to each child. This scale, which is sometimes called the "Oxford scale", has been first proposed by the OECD in 1982. We also probe the robustness of the results to the chosen scale in Appendix B, where we apply a different

¹In previous versions of the current paper we included older PSID waves and performed a consumption imputation to deal with the missing consumption categories, see De Giorgi et al. [2020].

equivalence scale. Of course the equivalence scales take into account the household size and composition.

Our measure of real adjusted family income (TFA) is

$$TFA_i = \left(\frac{Nominal \ Family \ Income \times 100}{CPI \times AEscale}\right).$$
(2)

We multiply Family income by 100 to preserve the scale of the variable given that CPI is equal to 100 in the base year. Similarly, we construct the following measure of real adjusted wealth as follows:

$$RealAdjWealth_i = \left(\frac{Nominal Wealth \times 100}{CPI \times AEscale}\right).$$
(3)

2.3 Household Consumption

Starting in 1999, the PSID asked its respondents the amount spent over different categories of goods, always at the household level. The expenditure categories for which we have information are the following: food at home, food out, food stamps (if used), rent, home insurance, electricity, heating, water, other utilities, car insurance, car repairs, gas, parking, bus, train, cab, other transportation, cost of school, cost of childcare, health insurance, expenditures on hospitals, doctors, and drugs. We construct our measure of actual consumption as the sum of all these expenditure categories.² This definition of total household consumption is the same as in Attanasio and Pistaferri [2014, 2016].³ Then, similarly to what we did already in the case of total family income, we deflate this measure by CPI and we divide it by the same equivalence scale as above, in order to take adequately into account different family compositions. Real adjusted family consumption is hence defined as

$$TC_i = \left(\frac{Nominal \ Family \ Consumption \times 100}{CPI \times AEscale}\right). \tag{4}$$

We multiply Family consumption by 100 to preserve the scale of the variable given that CPI is equal to 100 in the base year.

²Our consumption measure is constructed as comprehensively as possible given the information available in the PSID. However, some, limited, expenditure cathegories are not covered by the survey, such as holidays and investments in durables.

³Unfortunately, the measurement of consumption through surveys is bound to be plagued by error, this is extensively discussed for both the PSID and the CEX in [Carroll et al., 2015], we make no strides on that in the current paper. We note, however, that we are not interested in consumption level per se in the current work, but rather in the transition overtime.

	Whites				Blacks					
	Q1	Q_2	Q3	Q4	Q_5	Q1	Q2	Q3	Q4	Q5
Consumption	1819.77	3157.24	4374.82	5946.47	10653.04	1439.01	3115.53	4345.85	5943.8	9659.08
Expenditure	(568.88)	(367.84)	(361.20)	(576.22)	(4594.86)	(676.73)	(370.78)	(378.38)	(599.76)	(5078.71)
Total Family	6218.05	11081.72	14557.92	18986.3	28673.8	4691.11	7684.35	10549.85	12521.29	17335.16
Income	(10852.34)	(8434.98)	(9124.11)	(13532.52)	(31816.35)	(5663.31)	(4685.23)	(6606.61)	(7202.83)	(11174.41)
Family size	3.60	3.73	3.13	3.03	2.61	4.43	3.89	3.69	3.22	2.46
	(1.86)	(1.67)	(1.39)	(1.40)	(1.27)	(2.05)	(1.80)	(1.50)	(1.42)	(1.24)
Age	38.61	40.72	40.25	41.09	43.29	39.10	42.67	40.29	39.66	38.91
	(16.47)	(14.23)	(13.40)	(13.03)	(13.55)	(12.37)	(13.35)	(11.07)	(11.07)	(11.56)
Wealth	7644.13	22663.09	33935.34	53231.49	110725.4	1733.39	7608.48	9223.21	17930.7	50968.3
	(25221.23)	(54797.44)	(100764)	(233931.3)	(389992)	(6229.25)	(19273.47)	(17616.55)	(48726.1)	(213315.8)
Female	0.37	0.22	0.21	0.17	0.17	0.65	0.47	0.39	0.30	0.34
Education										
Grades 0-11	0.41	0.28	0.18	0.13	0.09	0.46	0.31	0.23	0.18	0.14
High School	0.39	0.38	0.34	0.31	0.24	0.4	0.4	0.4	0.38	0.28
Some College	0.16	0.17	0.26	0.26	0.25	0.13	0.24	0.31	0.27	0.37
BA or higher	0.04	0.18	0.22	0.3	0.42	0.01	0.05	0.06	0.17	0.21
Obs	1119	2348	2581	3110	3342	3050	1821	1587	1062	822

Table 1: Descriptive Statistics in base year 1999, by race and actual consumption quintile. Consumption expenditure is our measure of total consumption as expressed by eq. (4), Total Family Income is the variable defined by eq. (2), family size is the number of individuals in the family, age is age of the household head, wealth is real adjusted wealth as expressed by eq. (3), female is a dummy taking value 1 if the household head is a women and zero otherwise, and Education is one of the four education dummies standing, respectively, for grades 0-11, high school, some college, BA or higher degree.

In Table 1 we report the descriptive statistics for the main variables used in the present paper in the base year 1999. There are a few facts of interest in the simple descriptives. Consumption expenditure per adult equivalent ranges between about 1,400 - 1,800 USD (Q1 Blacks - Whites respectively) to almost 9,700 - 10,700 USD (Q5 Blacks - Whites respectively), with the two distributions being fairly close to each other in particular for the middle 3 quintiles. Total family income (per adult equivalent) ranges from 4,700 USD for Blacks in Q1 to almost 29,000 USD in Q5 for Whites with the Whites distribution dominating Blacks. Blacks tend to have larger families along the consumption distribution and in particular at the bottom of the distribution, i.e. 4.5 members vs. 3.6, but at the very top the situation is reversed with Blacks having smaller families than Whites (2.5 vs. 2.6 members). By using measures of total family income, total consumption and wealth that have been adjusted by an adult equivalence scale, as explained above, we are able to take into account differences in family composition between black and white households. Further, in Appendix B we probe the robustness of our results to the use of an alternative adult equivalence scale. Other two relevant facts stand out from Table 1 in terms of demographics: Blacks are younger than Whites at the top of the distribution by about 4.5 years. Perhaps, even more striking is the prevalence of female-headed (single parent) households among Blacks: 65% of households are female-headed in the bottom quintile vs. 37% for Whites, while at the top only 34% of Blacks households are female-headed (of course the share for Whites is lower at 17%). In terms of wealth it is noticeable how Whites have roughly three times larger wealth than Blacks along the distribution. Lastly, in terms of education it is easy to see how the large majority of households in the first quintile have at most a high-school degree (80% and 86% for Whites and Blacks respectively). We also note that in Q5 67% of Whites have at least some college, while that percentage is 58% for Blacks.

3 Unconditional Life-Cycle Consumption Dynamics

We turn now our attention to life-cycle dynamics. More specifically we describe racial differences in terms of dynamics over the life-cycle within the consumption distribution. We use rank-rank regressions to assess in which part of the distribution individuals starting from a given percentile end up ten years later. In Figure 1, each individual in a given year is assigned to a consumption percentile according to her position in the overall consumption distribution of that particular year (that is the consumption distribution including Blacks and Whites). Our aim here is to describe the dynamics along the national distribution rather than within race: we believe this to be the relevant measure of inequality.

In order to obtain insights on the differences in the degree of consumption persistence of Blacks and Whites along the overall distribution, we perform a rank-rank analysis in the spirit of Chetty et al. [2020]. We consider ten-year transitions, e.g. from 1999 to 2009, from 2001 to 2011 and so on. We then stack all the percentiles ranking in year t=1999, 2001,...,2007 and construct a second variable, ranking in t+10 so that we have all the transitions together in a stacked formulation. The unit of observation is individualby-year in 1999, 2001, ...,2007. We then run a local linear regression of $p_{t+10} = f(p_t)$, that smoother takes care of excess noise and non-parametrically produces the relationship between origin and destination ranking for Blacks and Whites separately. The rank-rank analysis allows us to address the issue of persistence over the life-cycle (10 years in this case), separately by race, with a simple graphic intuition of the results. We focus on 10 years transitions because these are long enough to give a peek into longer run mobility while preserving a reasonable number of observations, and limiting our ability to focus specific years start-end.

We perform this analysis separately by race in an unconditional fashion in Figure 1, later in the paper (Section 5) we will show how controlling for individual characteristics, namely age, gender, income, education and wealth changes the results. On the x-axis we have the origin percentiles of actual consumption while on the y-axis we report the destination percentiles after a 10-year span. In all graphs we include the 45 degree line indicating perfect persistence, i.e. an identical rank after 10 years. The flatter the slope of the rank-rank regressions, the higher is the mobility over the life-cycle. The blue line refers to Whites and the red line to Blacks. The figure reports the point estimate together with the 99% confidence bands.

We note several interesting facts. First, the blue line is always above the red line, meaning that for any possible percentile of origin in the consumption distribution, the average percentile of destination of Blacks is lower than that of Whites. The unconditional difference in the average destination percentile of consumption after 10 years between Blacks and Whites is around 10 percentiles at the top (80th percentile), and only slightly diminishes going toward the bottom of the consumption distribution. Overall, black individuals tend to shift downward in the consumption distribution. At the bottom of the distribution, white individuals tend to reach higher percentiles than black individuals, and at the top Whites are much more persistent than Blacks. Just to provide an example, from Figure 1 we see that if a Black individual was around the 99th percentile of consumption in year t, then on average she will end up below the 60th percentile after 10 years. While a White individual being in the 99th (top) percentile in year t will end up on average in the 68th percentile after 10 years. The vertical distance between the two lines relates to the rank distance of Bayer and Charles [2018]. Second, the intersection with the 45 degree line for Blacks coincide with about the 40th percentile while for Whites with the 55th. This means that 60% of Blacks are worse off than where they were ten years ago, while this is true only for about 45% of Whites.

In essence the rest of the paper is dedicated to understanding why these different life-cycle profiles emerge.



Figure 1: Average actual consumption rank after 10 years for an individual who was in each consumption percentile 10 years before, by race. The blue line refers to Whites and the red line to Blacks. 99% confidence bands are reported. Panel (a) no controls. PSID weights are used.

3.1 Transition matrices

As an exploratory analysis, below we report transition matrices between year t - 1 and year t for each available couple of years between 1999 and 2017, for consumption quintiles, separately for Blacks and Whites.

Transition matrices are one of the most commonly used methods to assess the degree of positional/rank mobility within an economy (see e.g. Shorrocks [1978], Fields and Ok [1999], Bonhomme and Robin [2009b]). In all the transition matrices presented in this Section, rows stand for past quintiles, whereas columns stand for present quintiles. The row totals of each matrix are equal to one. For example the cell in the second row, first column of Table 2, means that, among all individuals who were in the second consumption quintile in period t - 1, around 16% fell in the first (i.e. bottom) consumption quintile in period t. Conversely, the cell in the first row, second column of the same Table signifies that, among all individuals who were in the bottom quintile at time t - 1, around 24% ended up in the second quintile in period t. In Tables 2-5, an individual is recorded as a stayer if he/she is recorded to be in the same consumption quintile in period t - 1 and in period t, and a mover otherwise. The percentages of stayers are displayed in bold along the main diagonal of the transition matrix. Of course, a limitation of the transition matrix approach is that all intra-quintile transitions are disregarded.

Origin / Destination	1	2	3	4	5
1	0.4552	0.2066	0.1147	0.1097	0.1138
1	(0.0043)	(0.0035)	(0.0027)	(0.0027)	(0.0027)
9	0.1992	0.4643	0.2296	0.0841	0.0228
2	(0.0034)	(0.0042)	(0.0036)	(0.0024)	(0.0013)
3	0.0659	0.1844	0.4537	0.2256	0.0704
0	(0.0018)	(0.0028)	(0.0036)	(0.003)	(0.0018)
4	0.0382	0.0461	0.1932	0.5089	0.2135
4	(0.0013)	(0.0014)	(0.0026)	(0.0033)	(0.0027)
5	0.0258	0.0102	0.0433	0.1801	0.7406
5	(9.50E-04)	(6.10E-04)	(0.0012)	(0.0023)	(0.0026)

Table 2: Empirical 1-year transition matrices, Present quintile is on the columns, past quintile is on the rows. Row total is 1. Transitions are computed on the basis of total consumption (TC) for white individuals only.

Origin / Destination	1	2	3	4	5
1	0.6201	0.2538	0.0794	0.0326	0.0141
1	(0.0037)	(0.0033)	(0.0021)	(0.0013)	(9.00E-04)
0	0.2742	0.4658	0.1925	0.0544	0.013
2	(0.0038)	(0.0043)	(0.0034)	(0.0019)	(9.70E-04)
9	0.0819	0.2729	0.4265	0.1749	0.0439
3	(0.0028)	(0.0045)	(0.005)	(0.0039)	(0.0021)
4	0.0434	0.107	0.2585	0.4568	0.1343
4	(0.0026)	(0.0039)	(0.0055)	(0.0063)	(0.0043)
r.	0.0277	0.054	0.1073	0.2744	0.5367
J	(0.0031)	(0.0042)	(0.0058)	(0.0083)	(0.0093)

Table 3: Empirical 1-year transition matrices, Present quintile is on the columns, past quintile is on the rows. Row total is 1. Transitions are computed on the basis of total consumption (TC) for black individuals only.

The simple comparison of Table 2 and Table 3 provides descriptive evidence that Black individuals are less mobile than Whites at the bottom of the distribution, as the share of stayers in the bottom quintile is 62% vs around 46%. Further, Blacks are less persistent at the top of the consumption distribution, as the share of stayers is 54% vs around 74% for the Whites. Of course, this analysis is unconditional, i.e. it does not take into account the effect of covariates such as age, gender, income or education. For this reason, in what follows we present a similar analysis but in a conditional version, i.e. based on residual consumption quintiles, where residual consumption is estimated via the following

regression:

$$LTC_{it} = \alpha_0 + \alpha_1 Age_{it} + \alpha_2 AgeSq_{it} + \alpha_3 Educ_{it} + \alpha_4 Female_{it} + \alpha_5 LTFA_{it} + \varepsilon_{it}$$
(5)

where LTC_{it} is the log of total consumption as defined in eq. (4), Age_{it} and $AgeSq_{it}$ stand, respectively, for individual age and individual age squared, $Educ_{it}$ is the years of education achieved by the individual, $Female_{it}$ is a dummy equal to 1 if the individual is a women and zero otherwise and $LTFA_{it}$ is the log of total family income as defined in eq. (2). We then take the estimated residuals from this equation, $\hat{\varepsilon}_{it}$, and we construct the quintiles on the basis of this variables. In the following, we report transition matrices of the residuals between time t and time t - 1, for each available couple of years between 1999 and 2017, separately for black and white individuals. The results, reported in Tables 4 and 5, show that, even after controlling for the standard set of socio-demographic variables presented above, there are notable differences in the transition patterns of Blacks and Whites. Albeit such difference are now attenauted. Indeed, Blacks are still more persistent at the bottom of the distribution (61% vs 50% of stayers) and less persistent at the top of the residual consumption distribution (43% vs 58% of stayers in the top residual quintile).

Origin / Destination	1	2	3	4	5
1	0.4992	0.2399	0.1281	0.0771	0.0557
1	(0.0049)	(0.0041)	(0.0032)	(0.0026)	(0.0022)
2	0.1632	0.4179	0.2177	0.1287	0.0725
2	(0.0029)	(0.0039)	(0.0033)	(0.0026)	(0.0021)
3	0.0717	0.1991	0.3927	0.2194	0.1171
5	(0.0019)	(0.0029)	(0.0035)	(0.003)	(0.0023)
4	0.0417	0.0911	0.2174	0.4256	0.2242
7	(0.0014)	(0.002)	(0.0029)	(0.0035)	(0.0029)
5	0.0266	0.0588	0.101	0.2327	0.5809
0	(0.0011)	(0.0016)	(0.0021)	(0.0029)	(0.0034)

Table 4: Residual 1-year transition matrices, Present quintile is on the columns, past quintile is on the rows. Row total is 1. Transitions are computed on the basis of residual TC for white individuals only.

Origin / Destination	1	2	3	4	5
1	0.613	0.1952	0.0927	0.0534	0.0457
1	(0.004)	(0.0033)	(0.0024)	(0.0018)	(0.0017)
0	0.2698	0.3684	0.1863	0.105	0.0705
2	(0.0044)	(0.0047)	(0.0038)	(0.003)	(0.0025)
2	0.1637	0.244	0.3086	0.1771	0.1067
5	(0.0042)	(0.0048)	(0.0052)	(0.0043)	(0.0035)
4	0.1112	0.1478	0.2309	0.33	0.1801
4	(0.0039)	(0.0044)	(0.0053)	(0.0059)	(0.0048)
r.	0.1052	0.1187	0.1377	0.2102	0.4282
0	(0.0041)	(0.0043)	(0.0046)	(0.0055)	(0.0066)

Table 5: Residual 1-year transition matrices, Present quintile is on the columns, past quintile is on the rows. Row total is 1. Transitions are computed on the basis of residual TC for black individuals only.

The transition matrices reported above show descriptively the existence of a black/white differential in consumption persistence. This will be the main object of our investigation in the remainder of the paper.

4 The drivers of the Black/White consumption differential

4.1 Income

Let us analyze the variables that are most likely candidates to be the drivers of the Black/White consumption differential: income, education, and wealth. Let us now consider the Black/White life-cycle income profiles, across the consumption quintiles. In the following figures, each individual in a given year is assigned to an income quintile, which represents its position in the overall income distribution. Similarly, each individual is assigned to a consumption quintile on the basis of its consumption. Each individual may change quintile from one year to the following one, or he/she may remain in the same quintile. In Figure 2 we depict income profiles by race and by 1999 consumption quintiles. From this Figure we deduce that the total family income profile of the Whites is essentially always higher than that of the Blacks. The difference is smaller when individuals are young, i.e. in their 20s and until their 30s, then it increases and reaches its maximum when individuals are in their 50s. This is particularly true for individuals in the fourth and in the fifth top consumption quintiles.

In Figure 3 we report the histogram of total family income, divided by race and by 1999 consumption quintile, for 1999, our base year. From this Figure we deduce that, in all consumption quintiles, with possibly the exception of the bottom one, the density of total family income for the Whites (approximated by the histogram) lies to the right of that for the Blacks.



Figure 2: Life-cycle family income profile (log TFA is on the y-axis) for racial groups in the different 1999 consumption quintiles. Data for the 1999-2017 period. Panel (a) (bottom) first TC1999 quintile, panel (b) second TC1999 quintile, panel (c) third TC1999 quintile, panel (d) fourth TC1999 quintile, panel (e) (top) fifth TC1999 quintile.



Figure 3: Histogram of log Total Family Income (LTFA) by consumption quintile and race, in year 1999. Solid stands for Blacks, dashed for Whites. Panel (a) (bottom) first TC1999 quintile, panel (b) second TC1999 quintile, panel (c) third TC1999 quintile, panel (d) fourth TC1999 quintile, panel (e) (top) fifth TC1999 quintile.

4.2 Education

We now turn to document the education differentials in our data. The variable standing for education here can take four values, namely: (i) grades 0-11, (ii) high school, (iii) college dropout, (iv) college degree or higher.



Figure 4: Histogram of highest attained educational level by consumption quintile and race, in year 1999. Solid stands for Blacks, dashed for Whites. Panel (a) (bottom) first TC1999 quintile, panel (b) second TC1999 quintile, panel (c) third TC1999 quintile, panel (d) fourth TC1999 quintile, panel (e) (top) fifth TC1999 quintile. Educational levels are: (1) grades 0-11, (2) high school, (3) college dropout, (4) college degree or higher.

From Figure 4, we deduce that there are large differences in the highest educational level attained by race. This is a well-documented fact in the previous literature. In particular, from our data we deduce that the share of individuals with at least a college degree is notably higher among the Whites than among the Blacks in all consumption quintiles, with the difference becoming more and more evident in upper consumption quintiles. For example, in the fourth top consumption quintile this share is less than 20% for the Blacks, whereas it is around 30% for the Whites. Similarly, in the top fifth

consumption quintile, this share is more than 20% for the Blacks, but more than 40%, i.e. double, for the Whites.

4.3 Savings and Wealth

Finally, if wealth accumulation for Blacks is lower than for Whites, similar health and income shocks at the top of the consumption distribution will result in very different consumption movements. This is because Blacks are more exposed to the consequences of shocks due to a more limited buffer, and therefore the lack of self-insurance.⁴

Using an earlier sample, only composed by years 1984, 1989 and 1994, Gittleman and Wolff [2004] find no statistically significant difference in the stock of savings held by the Whites and Blacks. On the contrary, we document notable racial differences in wealth and stock of assets in our sample (1999-2017). Our results on the existence of a racial wealth gap are consistent with those by Altonji et al. [2000], who find that the racial gap in the wealth level cannot be fully explained by the distribution of income and demographic variables. However, the authors do not investigate the link of such a gap with differential consumption dynamics.

In Panel (a) of Figure 5, we plot the extensive margin for wealth holding, i.e. the percentage of individuals having positive wealth across consumption quintiles. The differences between Black and White households are 10-15 percentage points across all the consumption quintiles. From panel (b) it can be seen that, among people with a positive wealth, Blacks own far less of it than Whites, particularly at the top quintile. Wealth accumulated by Blacks is on average between 2 and 3 times smaller than that accumulated by the Whites in all the consumption quintiles. This difference is more evident at the top.

In Figure 5, we plot the extensive (left column) and intensive (right column) margins of total wealth (Panel (a) and (b)) and stock holding (Panel (c) and (d)) in the 5 quintiles of the 1999 consumption distribution. These variables are in real - per adult equivalent-terms. It is pretty easy to notice the large differences in both margins for Total Wealth at all quintiles. In the top quintile about 93% of Whites have positive wealth while this share is only 83% for Blacks; even more dramatic is the difference at the bottom quintile where about 79% of Whites have positive wealth with Blacks stopping at 63%. At the intensive

⁴As mentioned, we compute wealth as comprehensively as we can in the PSID, summing up seven asset types: imputed value of farm or business, imputed value of cash savings, imputed value of real estate other than home, imputed value of stocks, imputed value of vehicles, imputed value of other assets, value of home equity net of debt. This wealth measure is then divided by the Consumer Price Index (CPI), in order to obtain a measure of wealth in real terms. As above, we further divide this variable by the adult equivalence scale.







Figure 5: Total wealth: Panel (a) extensive margin: share of people having positive real wealth, by 1999 total consumption (TC) quintile (on the x-axis). Panel (b) intensive margin: real adjusted wealth by 1999 total consumption (TC) quintile (people with zero wealth have been excluded). Wealth is computed as the sum of seven asset types: imputed value of farm or business, imputed value of cash savings, imputed value of real estate other than home, imputed value of stocks, imputed value of vehicles, imputed value of other assets, value of home equity net of debt. Divided by the Consumer Price Index (CPI) and adult equivalence scale. Sample 1999-2017. Stock holdings: average value of stock holdings in US dollars divided by CPI and adult equivalent scale by 1999 TC quintile. Panel (c) extensive margin. Panel (d) intensive margin (people with no stocks have been excluded). Sample: 1999-2017. Top 1% of stocks has been trimmed in order to increase graph readability.

margin of total wealth, we see very large differences: in the bottom quintile Blacks hold on average around 8,000 USD, while Whites have slightly less than 17,000 USD (a ratio of 2.1); in the top quintile Blacks total only 41,000 USD of wealth while Whites are at about 148,000 USD (around 3.6 times as much).

The picture is even more striking if we consider stock holding (Panel (c) and (d)). Stocks are particularly interesting as they have provided about 7% returns in the last 20 years and could contribute disproportionately to the process of wealth accumulation. There are large racial differences in the percentage of households holding stocks. For example, in the top 1999 quintile, more than 40% of Whites households hold stocks, whereas only around 10% of Blacks households do. Second, there are also relevant differences in the amount of stocks held. In the top 1999 TC quintile, a White household on average 4,500 US dollars in stocks, whereas a Black one only holds around 2,500 US dollars. These differences in stock holdings are suggestive of large differential returns on assets and this difference is particularly relevant at the top consumption quintile.

Summing up, we find substantial racial differences in the amount of savings and wealth accumulation. In particular, the ratio of wealth held by Whites versus Blacks increases along the consumption distribution, and in general appears to be at least 2-3 times higher for Whites than for the Blacks. This is consistent with the results of the conditional rank-rank regressions that we will present in the next Section: Blacks are disproportionately exposed to downward mobility in the upper part of the distribution, but such downward mobility is muted once wealth is controlled for. Blacks save, on average, much less than Whites and accumulate much lower wealth. Hence, they are more prone to the risk of downward mobility across the consumption distribution.

5 Conditional Life-Cycle Consumption Dynamics

We turn now our attention the conditional version of the consumption dynamics described in Section 3.

We perform this analysis separately by race controlling for individual characteristics, namely age, gender, income, education and wealth. Age, gender and education are measured for the household head, whereas for income and wealth, as explained above, we use real adjusted measures that take into account the family composition, i.e. we divide by the adult equivalence scale. We compute wealth as comprehensively as we can in the PSID, summing up seven asset types: imputed value of farm or business, imputed value of cash savings, imputed value of real estate other than home, imputed value of stocks, imputed value of vehicles, imputed value of other assets, value of home equity net of debt. This wealth measure is then divided by the Consumer Price Index (CPI), in order to obtain a measure of wealth in real terms.

In performing this conditional rank-rank analysis, we use residual consumption for the analysis. In practice, first we regress actual consumption on the control variables as described above, then we rank the residuals of this regression for each year of the analysis. The approach allows then for the control variables to have different effects in different years.

In Figure 6, we report the results of the (smoothed) rank-rank conditional regressions. As in Figure 1, on the x-axis we have the origin percentiles of actual consumption while on the y-axis we report the destination percentiles after a 10-year span. In all graphs we include the 45 degree line indicating perfect persistence, i.e. an identical rank after 10 years. In the different panels of Figure 6 we use different sets of controls. Panel (a) displays the results for the unconditional case for ease of comparison, i.e. no controls. In Panel (b) the controls are age and gender; in Panel (c) the controls are age, gender, and total family income (TFA); in Panel (d) the controls are age, gender, and wealth. In Panel (e) the controls are age, gender and education. Finally, in Panel (f) the controls are age, gender, TFA, wealth and education.

Let us focus on the results when controlling for demographics and other factors, i.e. in Panels (b)-(f). We start by controlling for age and gender, as we showed in the descriptive table that these demographics appear to be important correlates of consumption rank. Panel (b) shows the results. The picture is similar to that for the unconditional case, with essentially no quantitative change in the destination percentile, neither for the Blacks, nor for the Whites. Despite the large difference in demographics shown in Table 1, those variables do not appear to close the gap between Blacks and Whites.



Figure 6: Average actual consumption rank in after 10 years for an individual who was in each consumption percentile in 10 years before, by race. The blue line refers to Whites and the red line to Blacks. 99% confidence bands are reported. Panel (a) no controls. Panel (b): controls: age and gender. Panel (c): controls as in (b) plus Total Family Income. Panel (d): controls as in (b) plus real adjusted wealth. Panel (e): controls as in (b) plus education. Panel (f): controls as in (b) plus education, TFA and real adjusted wealth. We control for wealth dynamically, e.g. we take the value of wealth in each of the years of the analysis. PSID weights are used in all the panels.

Another candidate to explain the differences in consumption dynamics is income. Consumption dynamics could differ simply because of racial differences in income dynamics. Thus, we repeat the analysis adding total family income (TFA) to the controls. Panel (c) reports the results. The differences between the percentiles of destination of Whites and Blacks are mitigated, but only at the top of the distribution, relative to the unconditional case. The difference in the average destination percentile of consumption is around 7 percentiles at the top (80th percentile), but stays at around 10 percentiles both at the 50th percentile and at the bottom (20th percentile). Income partially mitigates the gap between Blacks and Whites at the top but not at the bottom.

Further, wealth is crucial to insure away transitory and permanent shocks, while sustaining high consumption levels. We hence control for age, gender and wealth in Panel (d). The resulting picture is, not surprisingly, close to the one that we obtain when controlling for Total Family Income. In particular, in Panel (d) the gap in the average consumption destination percentile between Whites and Blacks is reduced (with respect to the unconditional case) to around 7-8 percentiles not only at the top, but across the whole consumption distribution.

In Panel (e), in addition to age and gender we also control for education. In this case, we notice that, unlike for the other controls, the gap is reduced to around 5-6 percentiles across the whole consumption distribution. The reduction is particularly evident at the top, where the gap is essentially closed down to 2-3 percentiles. Even if we are not able to control for the quality of education, we notice that the highest level of education achieved by the individual can explain a large part of the Black/White consumption gap, i.e. more than TFA and wealth can do.

Finally, In Panel (f) we control simultaneously for age, gender, TFA, real adjusted wealth and education. The gap is substantially reduced for all of the percentiles. It is essentially zero at the top of the consumption distribution and around 5 percentiles at the bottom and in the middle of the consumption distribution.

From the evidence above, we conclude that there are three factors that explain a substantial portion of life-cycle consumption racial differentials: income, wealth and education, However, while the three of them matter for the top of the distribution, at the bottom and the middle of the distribution education seems to be the most important variable to reduce the gap.

Note that the results presented in Figure 6 are such that, we first control for two variables that we deem fully exogenous (i.e. gender and age) and we then add a different additional covariate in each panel (i.e. TFA, education and real adjusted wealth). Only in the last panel all controls are included together. There is no sequential adding of covariates and we are able to evaluate the contributions of TFA, real adjusted wealth and education separately. Hence, we deem as not necessary to perform a decomposition in the spirit of Gelbach [2016].⁵



Figure 7: Impact of being Black on consumption percentile, based on the estimation of the difference between the actual TC percentiles of the Whites and the counterfactual TC percentiles of the Blacks. Blue stands for the Whites, red for the Blacks. Data for 1999-2017. Panel (a): we replace the actual distribution of the Blacks with their counterfactual TC percentiles distribution if they had the same age, gender and education distribution as the Whites. Panel (b): we replace the actual distribution of the Blacks with their counterfactual TC percentiles distribution if they had the same age, gender, education, log TFA and log adjusted real wealth distribution as the Whites. In Panel (b) top 1% of predicted consumption percentiles has been trimmed for graph readability. Panel (c): we replace the actual distribution of the Blacks with their counterfactual distribution and log TFA distribution as the Whites.

 $^{{}^{5}}$ Such a decomposition has been developed to analyze the contribution of different covariates to the race wage gap. Its framework is essentially static and does not fit well to our dynamic analysis of rank mobility.

The PSID also includes intergenerational information, such as the education level achieved by the parents of the individual. It would be possibile to include, say, father's education among the controls used for the analysis reported in the rank-rank regressions. However, father's education is likely to be highly correlated with both TFA, real adjusted wealth and individual educational level. This is why we do not include this additional variable in the previous analysis.

Our results reported in Figure 6 are robust to a series of robustness checks that are reported in Appendix B. The overall picture does not change if we use the Square Root Equivalence scale instead of the Oxford Equivalence. Further, our rank-rank regression results are robust to running the estimates on the SRC sample of the PSID only, as well as to using 5-year averages of TFA and real adjusted wealth (instead of annual values) as controls.

In Figure 7 we extend the rank-rank analysis to focus on the rank percentile gap, i.e. we investigate which percentile of the consumption distribution would the Blacks occupy if they had the same age, gender and education distribution than the Whites. This counterfactual exercise is carried out in the spirit of Bayer and Charles [2018]. In Panel (a) we notice a wide difference between the consumption rank distribution of the Whites and the counterfactual rank consumption distribution of the Blacks if they had the same distribution of age, gender and education as the Whites. From Panel (b) we notice that this difference in consumption ranks is substantially attenuated (but does not fully vanish) if we assign to the Blacks the same distribution of age, gender, education, TFA and real adjusted wealth as the Whites. These pictures are fully consistent with the results of our conditional rank-rank analysis reported above. Finally, in Panel (c) of Figure 7, we compare the distribution of real adjusted wealth for the Whites and the counterfactual distribution of real adjusted wealth for the Blacks, if they had the same age, gender, education and income distribution as the Whites. Taking the above-mentioned covariates into account, the difference between the two wealth distributions is attenuated, however it does not disappear.

6 Concluding Remarks

Our analysis strongly points towards the role of three main variables: income, education, and wealth as the key drivers of racial differentials in consumption dynamics. In particular, we show that Blacks at the top of the consumption distribution tend to fall in the ranking much more than Whites after a few years. At the same time while socio-demographics characteristics such as age or gender do not really close the life-cycle dynamics between Blacks and Whites in any part of the consumption distribution, income, education and wealth together, make the gap at the top of the distribution essentially disappear. Further, when controlling for all these three variables, the gap is essentially reduced (to around half of its original size) at the bottom and in the middle of the consumption distribution as well. It is well known, and confirmed in the current paper, that Blacks and Whites differ substantially in their amount of savings and wealth, it is however novel that we show how those differences persist even when comparing Blacks and Whites with initially similar levels of consumption.

References

- M. Aguiar and M. Bils. Has consumption inequality mirrored income inequality? American Economic Review, 105(9):2725–56, 2015.
- Joseph G. Altonji, Ulrich Doraszelski, and Lewis Segal. Black/white differences in wealth. *Economic Perspectives*, 24:1, 2000.
- A. B. Atkinson, T. Piketty, and E. Saez. Top incomes in the long run of history. *Journal* of economic literature, 49(1):3–71, 2011.
- O. Attanasio and L. Pistaferri. Consumption inequality over the last half century: some evidence using the new psid consumption measure. *American Economic Review*, 104 (5):122–26, 2014.
- O. Attanasio, E. Hurst, and L. Pistaferri. The evolution of income, consumption and leisure inequality in the united states 1980-2010. *Improving the measurement of consumer expenditures*, University of Chicago Press:100–140, 2014.
- O. P. Attanasio and L. Pistaferri. Consumption inequality. Journal of Economic Perspectives, 30(2):3–28, 2016.
- G. Auten, G. Gee, and N. Turner. Income inequality, mobility and turnover at the top in the us 1987-2010. *American Economic Review*, 103(3):168–72, 2013.
- D. H. Autor, L. F. Katz, and M. S. Kearney. Trends in us wage inequality: Revising the revisionists. *The Review of economics and statistics*, 90(2):300–323, 2008.
- P. Bayer and K. K. Charles. Divergent paths: A new perspective on earnings differences between black and white men since 1940. *The Quarterly Journal of Economics*, 133(3): 1459–1501, 2018.
- F. D. Blau and A. H. Beller. Black-white earnings over the 1970s and 1980s: Gender differences in trends. *The Review of Economics and Statistics*, pages 276–286, 1992.
- R. Blundell. Income dynamics and life-cycle inequality: Mechanisms and controversies. The Economic Journal, 124(576):289–318, 2014.
- R. Blundell and I. Preston. Consumption inequality and income uncertainty. *The Quarterly Journal of Economics*, 113(2):603–640, 1998.

- R. Blundell, L. Pistaferri, and I. Preston. Consumption inequality and partial insurance. American Economic Review, 98(5):1887–1921, 2008.
- R. Blundell, L. Pistaferri, and I. Saporta-Eksten. Consumption inequality and family labor supply. *American Economic Review*, 106(2):387–435, 2016.
- S. Bonhomme and J. M. Robin. Assessing the equalizing force of mobility using short panels: France, 1990–2000. The Review of Economic Studies, 76(1):63–92, 2009a.
- Stéphane Bonhomme and Jean-Marc Robin. Assessing the equalizing force of mobility using short panels: France, 1990–2000. The Review of Economic Studies, 76(1):63–92, 2009b.
- D. Card and A. B. Krueger. School quality and blacks-whites relative earnings: A direct assessment. *The Quarterly Journal of Economics*, 107(1):151–200, 1992.
- D. Card and A. B. Krueger. Trends in relative blacks-whites earnings revisited. The American Economic Review, 83(2):85–91, 1993.
- Christopher D. Carroll, Thomas F. Crossley, and John Sabelhaus. Improving the Measurement of Consumer Expenditures. University of Chicago Press, 06 2015. ISBN 9780226126654. doi: 10.7208/chicago/9780226194714.001.0001. URL https://doi.org/10.7208/chicago/9780226194714.001.0001.
- K. K. Charles, E. Hurst, and N. Roussanov. Conspicuous consumption and race. The Quarterly Journal of Economics, 124(2):425–467, 2009.
- K. Y. Chay and D. S. Lee. Changes in relative wages in the 1980s returns to observed and unobserved skills and blacks–whites wage differentials. *Journal of Econometrics*, 99(1): 1–38, 2000.
- R. Chetty, N. Hendren, P. Kline, and E. Saez. Where is the land of opportunity? the geography of intergenerational mobility in the united states. *The Quarterly Journal of Economics*, 129(4):1553–1623, 2014a.
- R. Chetty, N. Hendren, P. Kline, E. Saez, and N. Turner. Is the united states still a land of opportunity? recent trends in intergenerational mobility. *American Economic Review*, 104(5):141–47, 2014b.
- R. Chetty, N. Hendren, M. R. Jones, and S. R. Porter. Race and economic opportunity in the united states: An intergenerational perspective. *The Quarterly Journal of Economics*, 135(2):711–783, 2020.

- G. De Giorgi, L. Gambetti, and C. Naguib. Life-cycle inequality: Blacks and whites differentials in life expectancy, savings, income, and consumption. *CEPR DP*, (15182), 2020.
- Gary S Fields and Efe A Ok. The measurement of income mobility: an introduction to the literature. *Handbook of income inequality measurement*, pages 557–598, 1999.
- P. Ganong, D. Jones, P. Noel, D. Farrell, F. Greig, and C. Wheat. Wealth, race, and consumption smoothing of typical income shocks. Technical report, University of Chicago, Becker Friedman Institute for Economics Working Paper, (2020-49), 2020.
- Jonah B. Gelbach. When do covariates matter? and which ones, and how much? *Journal* of Labor Economics, 34(2):509–543, 2016. doi: 10.1086/683668.
- M. Gittleman and E. N. Wolff. Racial differences in patterns of wealth accumulation. Journal of Human Resources, 39(1):193–227, 2004.
- J. Heathcote, F. Perri, and G. L. Violante. Unequal we stand: An empirical analysis of economic inequality in the united states, 1967-2006. *Review of Economic dynamics*, 13 (1):15–51, 2010.
- J. J. Heckman, T. M. Lyons, and P. E. Todd. Understanding blacks-whites wage differentials, 1960-1990. American Economic Review, 90(2):344–349, 2000.
- J. S. Heywood and D. Parent. Performance pay and the whites-blacks wage gap. *Journal* of Labor Economics, 30(2):249–290, 2012.
- D. Krueger and F. Perri. Does income inequality lead to consumption inequality? evidence and theory. *The Review of Economic Studies*, 73(1):163–193, 2006.
- B. D. Meyer and J. X. Sullivan. Measuring the well-being of the poor using income and consumption. *Journal of Human Resources*, 38:4, 2003.
- R. L. Oaxaca and M. R. Ransom. On discrimination and the decomposition of wage differentials. *Journal of econometrics*, 61(1):5–21, 1994.
- J. Peoples and W. K. Talley. Black-whites earnings differentials: privatization versus deregulation. American Economic Review, 91(2):164–168, 2001.
- T. Piketty and E. Saez. Income inequality in the united states, 1913–1998. *The Quarterly journal of economics*, 118(1):1–41, 2003.

- G. E. Primiceri and T. Van Rens. Heterogeneous life-cycle profiles, income risk and consumption inequality. *Journal of monetary Economics*, 56(1):20–39, 2009.
- Anthony F Shorrocks. The measurement of mobility. *Econometrica: Journal of the Econometric Society*, pages 1013–1024, 1978.

Appendices

Appendix A Income and consumption profiles and dynamics

A.1 Life-cycle profiles

In all the results and Figures presented in this Section, each individual in a given year is assigned to an income quintile, which represents its position in the overall income distribution. Similarly, each individual is assigned to a consumption quintile on the basis of its consumption. Each individual may change quintile from one year to the following one, or he/she may remain in the same quintile. Our aim is precisely to determine whether there are relevant differences in quintile persistence between blacks and whites and if yes, in assessing whether these differences disappear once controlling for some observable covariates. In Figure B.1, we show the life-cycle consumption profile for Blacks and Whites in the five quintiles. The lines represent the average consumption for white and black individuals within the same quintile. This provides information about life-cycle evolution of consumption between the two racial groups.

It is apparent from the graphs that the consumption profile of the whites is essentially always higher than that of the blacks, in any of the considered quintiles. Differences are somehow smaller when the head of the household is young (i.e. in his 20s or 30s). On the contrary, this difference peaks when the age of the household head is equal to around 50 years old, especially in the bottom and in the top consumption quintiles.



Figure B.1: Life-cycle consumption profile for racial groups in the different total consumption (TC) quintiles. The reference year for the division into consumption quintiles is 1999. Data for the 1999-2017 period. Panel (a) (bottom) first TC1999 quintile, panel (b) second TC1999 quintile, panel (c) third TC1999 quintile, panel (d) fourth TC1999 quintile, panel (e) (top) fifth TC1999 quintile.

Figure B.2 depicts a similar profile. Indeed, also in the case of total family income, the life-cycle profile of the Whites is always higher than that of the Blacks, in any of the five quintiles. Again, the distance between the profiles of the two groups is smaller when the household head is young and then increases up to an age of the household head equal to 50-60 years old, especially in the top income quintile. Both income and consumption profiles display approximately the usual inverse-U shaped pattern in all the five quintiles, both for the Blacks and for the Whites. This shape is less evident in the bottom income and consumption quintiles. The income and consumption profiles reported in Figures B.1-B.2 are divided by the (consumption, resp. income) quintiles to which the individual belonged in 1999, our base year.



Figure B.2: Life-cycle family income profile for racial groups in the different TFA quintiles. The reference year for the division into consumption quintiles is 1999. Data for the 1999-2017 period. Panel (a) (bottom) first TFA1999 quintile, panel (b) second TFA1999 quintile, panel (c) third TFA1999 quintile, panel (d) fourth TFA1999 quintile, panel (e) (top) fifth TFA1999 quintile.



Figure B.3: Histogram of log TFA by TFA quintile and race, in year 1999. Solid stands for Blacks, dashed for Whites. Panel (a) (bottom) first TFA1999 quintile, panel (b) second TFA1999 quintile, panel (c) third TFA1999 quintile, panel (d) fourth TFA1999 quintile, panel (e) (top) fifth TFA1999 quintile.



Figure B.4: Histogram of log total consumption (TC) by TC quintile and race, in year 1999. Solid stands for Blacks, dashed for Whites. Panel (a) (bottom) first TC1999 quintile, panel (b) second TC1999 quintile, panel (c) third TC1999 quintile, panel (d) fourth TC1999 quintile, panel (e) (top) fifth TC1999 quintile.

Appendix B Robustness check

B.1 Equivalence scale

As a further check, we estimate persistence by adopting a different equivalence scale than the one in the main body of the paper. This means that we compute Total Family Income, Consumption and Wealth by dividing total family income/consumption/wealth by an alternative equivalence scale, i.e., the Square Root Scale. The formula applied here is the following:

$$SR = \sqrt{\text{Number of people in the household}}$$
 (6)

The results of the previous sections are confirmed, in the sense that the Black/White gap in consumption destination percentile is substantially reduced when both income, wealth and education are controlled for (as well as age and gender). The gap is essentially zero at the top of the consumption distribution, whereas 3-4 percentiles of difference remain at the bottom of it.



Figure B.5: Average actual consumption rank in after 10 years for an individual who was in each consumption percentile in 10 years before, by race. Panel (a) no controls. Panel (b): controls: age and gender. Panel (c): controls as in (b) plus Total Family Income. Panel (d): controls as in (b) plus real adjusted wealth. Panel (e): controls as in (b) plus education. Panel (f): controls as in (b) plus education, TFA and real adjusted wealth. We control for wealth dynamically, e.g. we take the value of wealth in each of the years of the analysis. 99% confidence intervals are reported. PSID weights are used in all the panels.

B.2 SRC sample only

Here we repeat the rank-rank analysis on the SRC random sample of the PSID only. SRC stands for Survey Research Center. PSID data used in the present paper include both the SRC and the SEO (Survey of Economic Opportunity) sample. This secold sample is included in the original data in order to have an oversampling of low-income households. The dataset with the SRC sample only for the years 1997-2017 consists of 85,883 Whites and 9,638 Blacks. Our main results are confirmed but the confidence intervals are wider (especially for the Blacks).



Figure B.6: Average actual consumption rank in after 10 years for an individual who was in each consumption percentile in 10 years before, by race. Panel (a) no controls. Panel (b): controls: age and gender. Panel (c): controls as in (b) plus Total Family Income. Panel (d): controls as in (b) plus real adjusted wealth. Panel (e): controls as in (b) plus education. Panel (f): controls as in (b) plus education, TFA and real adjusted wealth. We control for wealth dynamically, e.g. we take the value of wealth in each of the years of the analysis. 99% confidence intervals are reported. PSID weights are used in all the panels. SRC sample only.

B.3 Counterfactual exercise

In the following we perform the same analysis as in Figure 6 (rank-rank) by means of counterfactual exercises. The dependent variable is the difference in consumption percentiles between each year t - 10 and each year t in our sample.



Figure B.7: Outcome variable is the difference between consumption rank in year t-10 and in year t. We compare the consumption rank difference distribution of the Whites and the counterfactual rank difference distribution of the Blacks if they had the same distribution of certain covariates as the Whites. Such covariates include: Panel (a): constant only (no controls), Panel (b): age and gender, Panel (c): controls as in (b) plus Total Family Income. Panel (d): controls as in (b) plus real adjusted wealth. Panel (e): controls as in (b) plus education. Panel (f): controls as in (b) plus real adjusted wealth. Blue stands for the Whites, red for the Blacks.

B.4 5-year averages of TFA and real adjusted wealth

In the following we re-do the rank-rank analysis of Figure 6, but using 5-year averages of both real adjusted wealth and TFA as covariates (of course this slightly reduces the

sample size).



Figure B.8: Average actual consumption rank in after 10 years for an individual who was in each consumption percentile in 10 years before, by race. The blue line refers to Whites and the red line to Blacks. 99% confidence bands are reported. Panel (a) no controls. Panel (b): controls: age and gender. Panel (c): controls as in (b) plus 5-year average of Total Family Income. Panel (d): controls as in (b) plus 5-year average of real adjusted wealth. Panel (e): controls as in (b) plus education. Panel (f): controls as in (b) plus education, 5-year average of TFA and 5-year average of real adjusted wealth. PSID weights are used in all the panels.